

F2-1

Definition 2.2.2 A complex (real) matrix A is called unitary (orthogonal) iff $AA^{*T} = I$

Some facts about Hermitian symmetric matrices can be found in a textbook by R. Bellman, "Introduction to matrix analysis" Mc Graw Hill, 1970.

Theorem 2.2.1 If K is Hermitian symmetric then there exists a unitary matrix E such that

$$K = E \Lambda E^{*T}$$

where

$$\Lambda = \left[\begin{array}{ccc} \lambda_1 & & 0 \\ & \ddots & \\ 0 & & \lambda_n \end{array} \right]$$

with λ_i the eigenvalues of \underline{K} , not necessarily distinct.

In other words, Hermitian symmetric matrices are always diagonalizable

Theorem 2.2.2 A necessary and sufficient condition for such a K to be nonnegative definite is that $\lambda_i \geq 0 \quad \forall i = 1, \dots, n$

Theorem 2.2.3 Let K be Hermitian symmetric. Then for each distinct (simple) eigenvalue there corresponds an eigenvector which is orthogonal to all others. To each eigenvalue of multiplicity k there correspond k linearly independent eigenvectors, which are orthogonal to all eigenvectors of other eigenvalues.

However we can always perform a Gram-Schmit orthogonalization procedure and end up with k orthogonal eigenvectors. In summary, every Hermitian $(n \times n)$ matrix has n associated orthogonal eigenvalues $\{e_i\}_{i=1}^n$. In fact the matrix E of theorem 1 consists of these e_i 's as its columns. Returning to the factorization problem we want to find an $\underline{\mathrm{H}} \ \ni \ \underline{\mathrm{R}}_X =$

HH*T. We know that

$$\begin{array}{rcl}
\underline{R}_X & = & \underline{E}\underline{\Lambda}\underline{E}^{*T} \\
& = & \underline{E}\underline{\Lambda}^{1/2}\underline{\Lambda}^{1/2}\underline{E}^{*T}
\end{array}$$

where

$$\underline{\Lambda}^{1/2} \stackrel{\text{def}}{=} \begin{bmatrix} \sqrt{\lambda_1} & 0 & 0 & 0 \\ 0 & \sqrt{\lambda_2} & 0 & 0 \\ 0 & 0 & \ddots & 0 \\ 0 & 0 & 0 & \sqrt{\lambda_n} \end{bmatrix} = (\underline{\Lambda}^{1/2})^{*T}$$

$$\Rightarrow \underline{\mathbf{R}}_{X} = \underline{\mathbf{E}}\underline{\Lambda}^{1/2}\underline{\Lambda}^{1/2^{*T}}\underline{\mathbf{E}}^{*T}$$
$$= (\underline{\mathbf{E}}\underline{\Lambda}^{1/2})(\underline{\mathbf{E}}\underline{\Lambda}^{1/2})^{*T}$$

i.e. we arrived a t solution where

$$\underline{H}=\underline{E}\underline{\Lambda}^{1/2}$$

There exists the question of whether this solution is unique or not. The answer is no. To see this take any unitary matrix $\underline{U} \ni \underline{U}\underline{U}^{*T} = \underline{I}$. Then

$$\underline{\mathbf{R}}_{X} = (\underline{\mathbf{E}}\underline{\Lambda}^{1/2})\underline{\mathbf{I}}(\underline{\mathbf{E}}\underline{\Lambda}^{1/2})^{*^{T}} \\
= (\underline{\mathbf{E}}\underline{\Lambda}^{1/2}\underline{\mathbf{U}})(\underline{\mathbf{E}}\underline{\Lambda}^{1/2}\underline{\mathbf{U}})^{*^{T}} \\
= another\underline{\mathbf{H}}$$

Sometimes we take $\underline{\mathbf{U}} = \underline{\mathbf{E}}^{*T}$ and the resulting

$$\mathbf{H} = \mathbf{E} \mathbf{\Lambda}^{*T} \mathbf{E}^{*T}$$

is called the "square root" of R_X since then H is Hermitian symmetric. From an applications viewpoint this is useful in simulation, i.e. creating a random vector with desired correlation properties, starting from a "random number generator".

Note:

If $\underline{\mathbf{m}}_X \neq \underline{\mathbf{0}}$, then the appropriate linear transformation is

$$\bar{\mathbf{X}} = \bar{\mathbf{H}}\bar{\mathbf{W}} + \bar{\mathbf{m}}_X$$

where the factorization is done on \underline{K}_X , not \underline{R}_X .

Example:

Given the covariance matrix

$$\mathbf{K}_{X} = \left[\begin{array}{ccc} 1 & -1/2 & -1/2 \\ -1/2 & 1 & -1/2 \\ -1/2 & -1/2 & 1 \end{array} \right]$$

we find the eigenvalues by solving

$$det(\underline{\mathbf{K}}_X - \lambda_i \underline{\mathbf{I}}) = 0$$

$$\Rightarrow \lambda_1 = 0, \lambda_2 = \lambda_3 = 3/2$$

Solving for the eigenvectors we get

$$\lambda_1 = 0 \implies \underline{e}_1 = \frac{1}{\sqrt{3}} \begin{bmatrix} 1\\1\\1 \end{bmatrix}$$

$$\lambda_2 = 3/2 \implies \underline{e}_2 = \frac{1}{\sqrt{2}} \begin{bmatrix} 1\\-1\\0 \end{bmatrix}$$

$$\underline{e}_3 = \sqrt{\frac{2}{3}} \begin{bmatrix} 1/2\\1/2\\-1 \end{bmatrix}$$

So we could choose a linear transformation

$$\begin{split} \mathbf{H} &= \mathbf{E} \underline{\Lambda}^{1/2} &= \left[\mathbf{e}_1 \mid \mathbf{e}_2 \mid \mathbf{e}_3 \right] \begin{bmatrix} 0 & 0 \\ \sqrt{3/2} & \\ 0 & \sqrt{3/2} \end{bmatrix} \\ &= \begin{bmatrix} 0 & \sqrt{3/2} & 1/2 \\ 0 & -\sqrt{3/2} & 1/2 \\ 0 & 0 & -1 \end{bmatrix} \\ \Rightarrow \mathbf{X} &= \mathbf{H} \mathbf{W} = \begin{bmatrix} 0 & \sqrt{3/2} & 1/2 \\ 0 & -\sqrt{3/2} & 1/2 \\ 0 & 0 & -1 \end{bmatrix} \begin{bmatrix} W(u,1) \\ W(u,2) \\ W(u,3) \end{bmatrix} \end{split}$$

Notice that \underline{X} does not depend on W(u,1). In the above we solved the problem of spectral shaping which is equivalent to a covariance matrix factorization. The solution was unconstrained i.e. we imposed no restrictions on the nature of the linear transformation \underline{H} . We can pose an associated question: Can we find a causal matrix \underline{H} for the same factorization job?

Definition 2.2.3 In this context, causal means lower triangular i.e.

$$\begin{bmatrix} X(u,1) \\ \vdots \\ X(u,2) \end{bmatrix} = \begin{bmatrix} h_{11} & 0 & 0 \\ h_{21} & h_{22} & \\ \vdots & & \ddots & \\ h_{n1} & & \cdots & H_{nn} \end{bmatrix} \begin{bmatrix} W(u,1) \\ \vdots \\ W(u,n) \end{bmatrix} + \underline{\mathbf{m}}_{X}$$

or

$$X(u,i) = \sum_{i=1}^{i} h_{ij} W(u,j)$$

This is called the Cholesky Decomposition of positive definite matrices. We restate the problem as follows: Find a lower-triangular matrix \underline{H} such that

$$\underline{K}_{X} = \underline{H}\underline{H}^{*T}$$

Example:

For the real case

$$\begin{bmatrix} k_{11} & k_{12} & \cdots & k_{1n} \\ k_{12} & \cdot & & & \\ \vdots & & \ddots & & \\ k_{1n} & & \cdots & k_{nn} \end{bmatrix} = \begin{bmatrix} h_{11} & 0 & \cdots & 0 \\ h_{21} & h_{22} & & & \\ \vdots & & \ddots & & \\ h_{n1} & & \cdots & h_{nn} \end{bmatrix} \begin{bmatrix} h_{11} & h_{21} & \cdots & h_{n1} \\ 0 & h_{22} & & & \\ 0 & & \ddots & \vdots \\ 0 & & & h_{nn} \end{bmatrix}$$

From $h_{11}^2 = k_{11} \implies h_{11} = \pm \sqrt{k_{11}}$

$$\Rightarrow k_{12} = h_{21}h_{11} \Rightarrow h_{21} = \frac{k_{12}}{h_{11}}$$

In the same manner we can find the rest of the h_{ij} .

Using the concept of covariance factorization we can derive a set of insightful properties:

i. Spectral resolution

Assume a real covariance matrix \underline{K}_X . From the theory we know that \underline{K}_X can be decomposed as

$$\underline{\mathbf{K}}_{X} = \underline{\mathbf{E}}\underline{\mathbf{\Lambda}}\underline{\mathbf{E}}^{T}$$

where

$$\underline{\mathbf{E}} = [\underline{\mathbf{e}}_1 \mid \underline{\mathbf{e}}_2 \mid \cdots \mid \underline{\mathbf{e}}_n]$$

is the matrix of orthonormal eigenvectors of K_X and

$$\underline{\Lambda} = \left[\begin{array}{ccc} \lambda_1 & & 0 \\ & \ddots & \\ 0 & & \lambda_n \end{array} \right]$$

is the diagonal matrix of nonnegative eigenvalues. We can rewrite this as

$$\underline{\mathbf{K}}_{X} = \begin{bmatrix} \lambda_{1}\underline{\mathbf{e}}_{1} \mid \lambda_{2}\underline{\mathbf{e}}_{2} \mid \cdots \mid \lambda_{n}\underline{\mathbf{e}}_{n} \end{bmatrix} \begin{bmatrix} \underline{\mathbf{e}}_{1}^{T} \\ --\\ \underline{\mathbf{e}}_{2}^{T} \\ \vdots \\ --\\ \underline{\mathbf{e}}_{n}^{T} \end{bmatrix}$$

or

$$\underline{\mathbf{K}}_{X} = \sum_{i=1}^{n} \lambda_{i} \underline{\mathbf{e}}_{i} \underline{\mathbf{e}}_{i}^{T}$$

This shows that K_X can be decomposed (resolved) into a sum of n matrices, each of the form $\underline{e}_i\underline{e}_i^T$, with weight λ_i . The set of n eigenvectors $\{\underline{e}_i\}_{i=1}^n$ constitutes n basis for the n-dimensional vector space and each deterministic vector \underline{A} can be expanded into a series

$$\underline{\mathbf{A}} = \sum_{i=1}^n a_i \underline{\mathbf{e}}_i$$

where

$$a_i = (\underline{\mathbf{A}}, \underline{\mathbf{e}}_i) = \underline{\mathbf{A}}^T \underline{\mathbf{e}}_i$$

So far we have determined that given some covariance matrix K_X , we can find its n orthonormal eigenvectors \underline{e}_i and use them as a basis of an n-dimensional vector space. Each deterministic vector \underline{A} can be described in terms of its "projections" a_j along the \underline{e}_j coordinate. It is also clear that we can create random vectors by choosing these projections to be random variables $A_i(u)$, i.e.

$$\underline{\mathbf{A}} = \sum_{i=1}^{n} A_i(u)\underline{\mathbf{e}}_i$$

Note: If the eigenvectors have the form $\underline{e}_i = [0 \cdots 1 \cdots 0]^T$ with 1 in the i^{th} position then

$$\underline{\mathbf{A}} = \left[\begin{array}{c} A_1(u) \\ \vdots \\ A_n(u) \end{array} \right]$$

The question arises whether these random coefficients $A_i(u)$ can be chosen in such a way that the resulting \underline{A} is actually $\underline{X}(u)$ i.e. the vector whose covariance is the given one \underline{K}_X . We will answer that later (whitening). Note that whitening is the reverse problem to "coloring" or "spectral shaping".